Localisation-Resistant Random Words with Small Alphabets

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Introduction

Object of study: Probability distributions on (some subset of) $\{1, \ldots, q\}^n$ and $\{0, 1\}^n$.

Examples

q-coloured words:
$$\mathcal{C}_{q,n} = \{x_1 \dots x_n \in \{1,\dots,q\}^n \mid x_i \neq x_{i+1}\}.$$

Independent-set words:
$$J_n = \{y_1 \dots y_n \in \{0,1\}^n \mid y_i = 1 \Rightarrow y_{i-1} = y_{i+1} = 0\}.$$

Localisation resistance: No information is revealed about location of subword(s).

+	$\stackrel{\geq k}{\longleftrightarrow}$			
 X		У		
 X			у	
 У		X		

all occur with the same probability.

Introduction

Main results (informal)

- Complete description of localisation-resistant probability distributions on independent-set words.
- We need q ≥ 4 for n large enough in order for a localisation-resistant probability distribution on q-coloured words to exist.

Motivation

- *q*-colouring using local resources is a fundamental problem in Distributed Computing [Pettie et al.- STOC '18].
- Open question: are there faster algorithms if we have access to quantum resources?
- q-colouring using quantum resources must be localisation-resistant [Gavoille, Kosowki & Markiewicz '09].

Some definitions

 $x_1 \dots x_n$ a random word on a small alphabet.

Stationarity: "Shift invariance for probability distributions".

For all $I, J \subset \{1, ..., n\}$ intervals at distance at least k of each other:

$$k$$
-dependence: $Pr(x_I, x_J) = Pr(x_I) \cdot Pr(x_J)$. k -localisability: $Pr(x_I, x_J)$ depends only on $|I|$, $|J|$.

Facts

- 0-dependence = independence.
- 0-localisability = exchangeability.
- k-localisability ⇒ stationarity.
- k-dependence + stationarity ⇒ k-localisability.



k-dependence VS k-localisability

k-dependence + stationarity $\Rightarrow k$ -localisability.

BUT

k-localisability $\not\Rightarrow m$ -dependence for some m.

Example

 $S = \{\text{all permutations of } \{1, \dots, n\}\}$

Pr: uniform distribution on S.

This distribution is 0-localisable since

$$Pr(x_I, x_J) = \frac{(n - |I| - |J|)!}{n!}$$

But not k-dependent for any $k \le n$ because

$$\Pr(x_I) \cdot \Pr(x_J) = \frac{(n-|I|)!}{n!} \cdot \frac{(n-|J|)!}{n!} \neq \Pr(x_I, x_J)$$



The *q*-colouring problem

Recall that the set of *q*-coloured words of length *n* is:

$$C_{q,n} = \{x_1 \dots x_n \in \{1, \dots, q\}^n \mid x_i \neq x_{i+1}\}$$

1-dependent colouring

[Holroyd & Liggett '15, '16]:

- There is a stationary 1-dependent *q*-colouring for all $q \ge 4$ for every $n \in \mathbb{N}$.
- There is no stationary 1-dependent 3-colouring for *n* large enough.

Because "k-dependent + stationary $\Rightarrow k$ -localisable", this implies

1-localisable colouring

There is a 1-localisable *q*-colouring for all $q \ge 4$ for every $n \in \mathbb{N}$.



The *q*-colouring problem

1-localisable colouring

There is a 1-localisable *q*-colouring for all $q \ge 4$ for every $n \in \mathbb{N}$.

Is 1-localisable 3-colouring possible?

Main result

Every 1-localisable probability distribution for random q-coloured words of length n requires $q \ge 4$ for n large enough.

q-coloured words and independent-set words

Recall

q-coloured words: $\mathcal{C}_{q,n} = \{x_1 \dots x_n \in \{1,\dots,q\}^n \mid x_i \neq x_{i+1}\}.$ Independent-set words: $\mathfrak{I}_n = \{y_1 \dots y_n \in \{0,1\}^n \mid y_i = 1 \Rightarrow y_{i-1} = y_{i+1} = 0\}.$

Choose any colour c and define a function $f = (f_1, \dots, f_n)$ on $\mathcal{C}_{q,n}$ as follows:

$$y_i = f_i(x_i) = 1$$
 if $x_i = c$ and $y_i = f_i(x_i) = 0$ if $x_i \neq c$

Then

- The image of f is exactly \mathfrak{I}_n .
- *f* preserves *k*-dependence, *k*-localisability, stationarity.
- $\Pr(y_i = 1) \ge 1/q$.

Goal

Find $\max \Pr(y_i = 1)$ for a 1-localisable probability distribution on \mathfrak{I}_n .



We aim to find max $Pr(y_i = 1)$ for a 1-localisable probability distribution on \mathfrak{I}_4 . Let $p_1 = Pr(y_i = 1)$, $p_2 = Pr(y_i = 1, y_{i+2} = 1)$.

<i>Y</i> 1 <i>Y</i> 2 <i>Y</i> 3 <i>Y</i> 4	$Pr(y_1y_2y_3y_4)$
0101	
1010	
1001	
0100	
0010	
1000	
0001	
0000	

$$Pr(1010) = Pr(1010) = Pr(1001) = p_2.$$

We aim to find max $Pr(y_i = 1)$ for a 1-localisable probability distribution on \mathfrak{I}_4 . Let $p_1 = Pr(y_i = 1)$, $p_2 = Pr(y_i = 1, y_{i+2} = 1)$.

<i>Y</i> 1 <i>Y</i> 2 <i>Y</i> 3 <i>Y</i> 4	$Pr(y_1y_2y_3y_4)$
0101	p_2
1010	p_2
1001	p_2
0100	
0010	
1000	
0001	
0000	

$$Pr(0100) + p_2 = Pr(0100) + Pr(0101) = Pr(010*) = p_1.$$

We aim to find max $Pr(y_i = 1)$ for a 1-localisable probability distribution on \mathcal{I}_4 . Let $p_1 = Pr(y_i = 1)$, $p_2 = Pr(y_i = 1, y_{i+2} = 1)$.

<i>Y</i> 1 <i>Y</i> 2 <i>Y</i> 3 <i>Y</i> 4	$Pr(y_1y_2y_3y_4)$
0101	p_2
1010	p_2
1001	p_2
0100	$p_1 - p_2$
0010	$p_1 - p_2$
1000	
0001	
0000	

$$Pr(1000) + 2p_2 = Pr(1000) + Pr(1010) + Pr(1001) = Pr(10 \star \star) = p_1.$$

We aim to find max $Pr(y_i = 1)$ for a 1-localisable probability distribution on \mathcal{I}_4 . Let $p_1 = Pr(y_i = 1)$, $p_2 = Pr(y_i = 1, y_{i+2} = 1)$.

<i>Y</i> 1 <i>Y</i> 2 <i>Y</i> 3 <i>Y</i> 4	$\Pr(y_1y_2y_3y_4)$
0101	p_2
1010	p_2
1001	p_2
0100	$p_1 - p_2$
0010	$p_1 - p_2$
1000	$p_1 - 2p_2$
0001	$p_1 - 2p_2$
0000	

Pr(0000) = 1 minus the rest.

We aim to find max $Pr(y_i = 1)$ for a 1-localisable probability distribution on \mathfrak{I}_4 . Let $p_1 = Pr(y_i = 1)$, $p_2 = Pr(y_i = 1, y_{i+2} = 1)$.

<i>Y</i> 1 <i>Y</i> 2 <i>Y</i> 3 <i>Y</i> 4	$Pr(y_1y_2y_3y_4)$
0101	p_2
1010	p_2
1001	p_2
0100	$p_1 - p_2$
0010	$p_1 - p_2$
1000	$p_1 - 2p_2$
0001	$p_1 - 2p_2$
0000	$1-4p_1+3p_2$

We aim to find max $Pr(y_i = 1)$ for a 1-localisable probability distribution on \mathfrak{I}_4 . Let $p_1 = Pr(y_i = 1)$, $p_2 = Pr(y_i = 1, y_{i+2} = 1)$.

<i>Y</i> 1 <i>Y</i> 2 <i>Y</i> 3 <i>Y</i> 4	$Pr(y_1y_2y_3y_4)$
0101	p_2
1010	p_2
1001	p_2
0100	$p_1 - p_2$
0010	$p_1 - p_2$
1000	$p_1 - 2p_2$
0001	$p_1 - 2p_2$
0000	$1-4p_1+3p_2$

Solving $p_i \ge 0$ and $\Pr(y_1y_2y_3y_4) \ge 0$ we find the maximum value to be $p_1 = 2/5$.

General case

Set
$$p_1 := \Pr(1 \star \cdots \star), p_2 := \Pr(1 \star 1 \star \cdots \star), \ldots, p_\ell$$
.

Find maximum value of p_1 such that $p_i \ge 0$ and $\Pr(y_1 \dots y_n) \ge 0$ (where $n = 2\ell - 1$ or $n = 2\ell$).

- Each $\Pr(y_1 \dots y_n) \ge 0$ is uniquely determined as a linear function of p_1, \dots, p_ℓ .
- We thus need to solve a linear programming problem.
- **Problem**: Exponential (in ℓ) number of constraints.

Equivalent subsystem

$$\begin{aligned} & \Pr(0^n) \geq 0 \\ & \Pr((10)0^{n-2} \geq 0 \\ & \Pr((10)^20^{n-4}) \geq 0 \\ & \dots \\ & \Pr((10)^{\ell-1}0^{n-2\ell+2)} \geq 0 \end{aligned}$$

 \Rightarrow Linear programming problem in ℓ constraints and ℓ variables.



Linear programming

Formulation using matrices and vectors

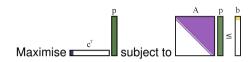
Maximise
$$\mathbf{c}^{\mathsf{T}}\mathbf{p}(=p_1)$$
, subject to $\mathbf{A}\mathbf{p} \leq \mathbf{b}$ and $\mathbf{p} \geq \mathbf{0}$.

Here the matrix and the vectors are of the form:

$$\mathbf{A} = egin{bmatrix} a_{1,1} & a_{1,2} & \dots & \dots & a_{1,\ell-1} & a_{1,\ell} \ -1 & a_{2,2} & \dots & \dots & a_{2,\ell-1} & a_{2,\ell} \ 0 & -1 & \ddots & \ddots & \vdots & \vdots \ \vdots & \vdots & \ddots & \ddots & \vdots & \vdots \ 0 & 0 & \dots & -1 & a_{\ell-1,\ell-1} & a_{\ell-1,\ell} \ 0 & 0 & \dots & 0 & -1 & a_{\ell,\ell} \end{bmatrix}$$

$$a_{i,j} = (-1)^{i+j} {2\ell+2-(i+j) \choose j-i+1}, \mathbf{c} = (1,0,\ldots,0), \mathbf{b} = (1,0,\ldots,0), \mathbf{p} = (p_1,\ldots,p_\ell).$$

Linear programming



Theorem

For linear programming problems that have the general form above, the optimal value for $\mathbf{c}^\mathsf{T}\mathbf{p}$ (= p_1) is $\frac{u_\ell}{u_{\ell+1}}$ where the sequence $(u_k)_{k\geq 1}$ is defined by $u_1=1$ and

$$u_{k+1} = \sum_{i=1}^{k} a_{\ell-k+1,\ell-k+i} u_{k+1-i}$$

This optimal value is obtained by solving the special case Ap = b.

We prove this theorem using the duality theorem for linear programming.



Linear programming

Applying the theorem on the previous slide to our problem $(a_{i,j} = (-1)^{i+j} {2\ell+2-(i+j) \choose j-i+1})$ we get

Theorem

The optimal value for p_1 is $\frac{u_\ell}{u_{\ell+1}}$ where $u_1=1$ and $u_{k+1}=\sum_{i=1}^k (-1)^{i+1} \binom{2k+1-i}{i} u_{k+1-i}$.

This is exactly the sequence of **Catalan numbers** $(c_n)_{n\in\mathbb{N}}$, $c_n = \frac{1}{n+1}\binom{2n}{n}$.

Finally

$$\frac{c_\ell}{c_{\ell+1}} = (\ell+2)/(4\ell+2)$$
, hence the optimal value of $p_1 \to 1/4$ as $\ell \to \infty$.



Further results and questions

- 1. (In paper) If we plug in a feasible value for p_1 , we get a linear programming problem for maximising p_2 that has the same general form and we may apply the same theorem and so on for p_3 , etc. This way we get $p_i \le (c_{\ell-i+1}/c_{\ell+1}) \to 1/4^i$ as $\ell \to \infty$.
 - 2. (Proven, not in paper) Let

$$\mathfrak{I}_{k,n} = \{y_1 \dots y_n \in \{0,1\}^n \mid y_i = 1 \Rightarrow y_{i-j} = y_{i+j} = 0, 1 \le j \le k\}$$

We have a similar result for k-localisable probability distributions on $\mathfrak{I}_{k,n}$ with the sequence of **Fuss-Catalan numbers**.

3. **(Open questions)** What about combinatorial structures other than words? E.g. labellings on graphs?

